

## CHI-SQUARED TESTS FOR CATEGORICAL DATA

In a **multinomial** experiment, the independent experimental units are classified to one of  $k$  categories determined by the levels of a discrete factor. Let  $n_1, n_2, \dots, n_k$  be the counts of the numbers of experimental units in the  $k$  categories, where  $n_1 + n_2 + \dots + n_k = n$ .

The probability that an experimental unit is classified to category  $i$  is  $p_i$ , for  $i = 1, \dots, k$ , so that

$$p_1 + p_2 + \dots + p_k = 1.$$

- The **one-way** classification table can be displayed as follows:

Category	1	2	...	$k$
Count	$n_1$	$n_2$	...	$n_k$
Probability	$p_1$	$p_2$	...	$p_k$

We can test a hypothesis  $H_0$  that fully specifies  $p_1, \dots, p_k$ , for example

$$H_0 : p_1 = p_1^{(0)}, p_2 = p_2^{(0)}, \dots, p_k = p_k^{(0)}$$

so that, for  $k = 3$ , we might have

$$H_0 : p_1 = p_2 = p_3 = 1/3 \quad \text{or} \quad H_0 : p_1 = 1/2, p_2 = p_3 = 1/4.$$

We use the test statistic

$$X^2 = \sum_{i=1}^k \frac{(n_i - np_i^{(0)})^2}{np_i^{(0)}} = \sum_{i=1}^k \frac{(\text{Observed Count in Cell } i - \text{Expected Count in Cell } i)^2}{\text{Expected Count in Cell } i}$$

We sometimes write  $\hat{n}_i = np_i^{(0)}$ . If  $H_0$  is true,

$$X^2 \sim \text{Chi-squared}(k - 1).$$

- The **two-way** classification table can also be constructed to represent the cross-classification for two discrete factors  $A$  and  $B$  with  $r$  and  $c$  levels respectively.

		Factor B			
		1	2	...	$c$
Factor A	1	$n_{11}$	$n_{12}$	...	$n_{1c}$
	2	$n_{21}$	$n_{22}$	...	$n_{2c}$
	⋮	⋮	⋮		⋮
	$r$	$n_{r1}$	$n_{r2}$	...	$n_{rc}$

To test the hypothesis

$$H_0 : \text{Factor A and Factor B levels are assigned independently}$$

we use the same test statistic that can be rewritten

$$X^2 = \sum_{i=1}^r \sum_{j=1}^c \frac{(n_{ij} - \hat{n}_{ij})^2}{\hat{n}_{ij}}$$

where

$$\hat{n}_{ij} = \frac{n_{i.} n_{.j}}{n} \quad n_{i.} = \sum_{j=1}^c n_{ij} \quad n_{.j} = \sum_{i=1}^r n_{ij}.$$

The terms  $n_{i.}$  and  $n_{.j}$  are the row and column totals for row  $i$  and column  $j$  respectively. If  $H_0$  is true

$$X^2 \sim \text{Chi-squared}((r - 1)(c - 1))$$

**EXAMPLE 1: DNA Sequence Data**

The counts of the numbers of nucleotides (A,C,G,T) in the DNA sequence of the cancer-related gene BRCA 2 are presented in the table below.

Category	1	2	3	4	Total
Nucleotide	A	C	G	T	
Count	38514	24631	25685	38249	127079

so that  $k = 4$ . To test the hypothesis

$$H_0 : p_1 = p_2 = p_3 = p_4 = 1/4$$

We use the one-way table chi-squared test: here

$$\hat{n}_i = np_i^{(0)} = \frac{127079}{4} = 31769.75$$

so the test statistic is

$$\begin{aligned} X^2 &= \frac{(38514 - 31769.75)^2}{31769.75} + \frac{(24631 - 31769.75)^2}{31769.75} + \frac{(25685 - 31769.75)^2}{31769.75} + \frac{(38249 - 31769.75)^2}{31769.75} \\ &= 5522.597 \end{aligned}$$

We compare this with the Chi-squared( $k - 1$ )  $\equiv$  Chi-squared(3) distribution. From McClave and Sincich, p. 898,

$$\text{Chisq}_{0.05}(3) = 7.815 < X^2$$

so  $H_0$  is **rejected**.

**EXAMPLE 2: Eye and Hair Colour Data**

The table below contains counts of the number of people in a study with a combination of eye and hair colour.

		Hair				$n_{i.}$
		Black	Brunette	Red	Blonde	
Eyes	Brown	68	119	26	7	220
	Blue	20	84	17	94	215
	Hazel	15	54	14	10	93
	Green	5	29	14	16	64
	$n_{.j}$	108	286	71	127	592

so  $r = c = 4$ . To test the hypothesis

$$H_0 : \text{Eye and Hair colour are assigned independently}$$

we use the  $X^2$  statistic

$$X^2 = \sum_{i=1}^r \sum_{j=1}^c \frac{(n_{ij} - \hat{n}_{ij})^2}{\hat{n}_{ij}}$$

Here, for example, for  $i = 2$  and  $j = 3$

$$\hat{n}_{23} = \frac{n_{2.} \times n_{.3}}{n} = \frac{215 \times 71}{592} = 25.785.$$

In fact, on complete calculation, we find that

$$X^2 = 138.2898.$$

We compare this with the Chi-squared( $(r - 1)(c - 1)$ )  $\equiv$  Chi-squared(9) distribution. From McClave and Sincich, p. 898,

$$\text{Chisq}_{0.05}(9) = 16.919 < X^2$$

so  $H_0$  is **rejected**

## Chi-Squared test for the nucleotide count data

Use

*Analyze* → *Nonparametric Tests* → *Chi-Square*

pulldown menus.

For the test of

$$H_0 : p_1 = p_2 = p_3 = p_4 = 1/4$$

First null hypothesis

Nucleotide			
	Observed N	Expected N	Residual
A	38514	31769.8	6744.3
C	24631	31769.8	-7138.8
G	25685	31769.8	-6084.8
T	38249	31769.8	6479.3
Total	127079		

Chi-squared Statistic = 5522.597

### Test Statistics

	Nucleotide
Chi-Square(a)	5522.597
df	3
Asymp. Sig.	.000

p-value < 0.001

a 0 cells (.0%) have expected frequencies less than 5. The minimum expected cell frequency is 31769.8.

For the test of

$$H_0 : p_1 = p_4 = 0.3 \quad p_2 = p_3 = 0.2$$

Second null hypothesis

Nucleotide			
	Observed N	Expected N	Residual
A	38514	38123.7	390.3
C	24631	25415.8	-784.8
G	25685	25415.8	269.2
T	38249	38123.7	125.3
Total	127079		

Chi-squared Statistic = 31.492

### Test Statistics

	Nucleotide
Chi-Square(a)	31.492
df	3
Asymp. Sig.	.000

p-value < 0.001

a 0 cells (.0%) have expected frequencies less than 5. The minimum expected cell frequency is 25415.8.

## Chi-Squared test for the Hair and Eye colour count data

Use

*Analyze* → *Descriptive Statistics* → *Crosstabs*  
pulldown menus.

For the test of

$H_0$  : Hair and Eye colour are assigned independently

**Eye Colour \* Hair Colour Crosstabulation**

Count

		Hair Colour				Total
		Black	Brown	Red	Blond	
Eye Colour	Brown	68	119	26	7	220
	Blue	20	84	17	94	215
	Hazel	15	54	14	10	93
	Green	5	29	14	16	64
Total		108	286	71	127	592

**Chi-Square Tests**

	Value	df	Asymp. Sig. (2-sided)
Pearson Chi-Square	138.290(a)	9	.000
Likelihood Ratio	146.444	9	.000
Linear-by-Linear Association	28.292	1	.000
N of Valid Cases	592		

p-value < 0.001

a. 0 cells (.0%) have expected count less than 5. The minimum expected count is 7.68.

Chi-square statistic = 138.290

Note the comment returned by SPSS: The chi-squared test is not appropriate if any of the cells in the table have expected count less than 5 under the null hypothesis.

In this case, there is no problem as the cell counts are large enough.