

Use this page for extra answer or working space

Answer to Q1:

(a) The proposed model is

$$Y_i = \beta_0 + \beta_1 x_{i1} + \epsilon_i$$

so that the fitted value under the model is  $\beta_0 + \beta_1 x_{i1}$ . Therefore, the vertical mis-fit is  $y_i - (\beta_0 + \beta_1 x_{i1})$ , which on aggregating the squared values takes

$$S(\beta_0, \beta_1) = \sum_{i=1}^n (y_i - (\beta_0 + \beta_1 x_{i1}))^2.$$

Least squares estimation proceeds by minimizing this function with respect to the  $\beta$  parameters.

5 Marks

(b) The least squares function is, in vector form,

$$S(\beta) = (\mathbf{y} - \mathbf{X}\beta)^\top (\mathbf{y} - \mathbf{X}\beta)$$

which, on differentiation, yields

$$\frac{\partial S(\beta)}{\partial \beta} = -2\mathbf{X}^\top (\mathbf{y} - \mathbf{X}\beta).$$

Equating to zero and solving, we see that at the solution

$$\mathbf{X}^\top (\mathbf{y} - \mathbf{X}\hat{\beta}) = \mathbf{0}_2.$$

5 Marks

(c) The solution to the least squares estimation problem can be written

$$\hat{\beta} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y} = \mathbf{A} \mathbf{y}$$

say, which is a linear transform of  $\mathbf{y}$ . The fitted values are

$$\hat{\mathbf{y}} = \mathbf{X}\hat{\beta} = \mathbf{X}(\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y} = \mathbf{H} \mathbf{y}$$

which is also a linear combination.

4 Marks

(d) Here are the full results:

```
1 > summary(lm(y~x))
2 Coefficients:
3             Estimate Std. Error t value Pr(>|t|)
4 (Intercept)   1.4928      1.7799   0.839   0.4293
5 x             0.3635      0.1388   2.619   0.0345 *
6 ---
7 Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
8
9 Residual standard error: 1.704 on 7 degrees of freedom
10 Multiple R-squared:  0.4948,    Adjusted R-squared:  0.4227
11 F-statistic: 6.857 on 1 and 7 DF,  p-value: 0.03448
```

The answers are as follows:

- (i) the sample size  $n$ ;  $n = 9$  (from line 9) 1 Mark
- (ii) the `Estimate` omitted from line 4: 1.4928; 1 Mark
- (iii) the conclusion of the test of the null hypothesis  $H_0 : \beta_1 = 0$  omitted from line 5: reject the null hypothesis as the test statistic is significant (note the  $p$ -value on line 11 which is identical to the  $p$ -value in the test referred to). 1 Mark
- (iv) whether  $x$  is a useful predictor of  $y$ : it is useful, based on the test result above and the  $R^2$  statistic on line 10. 1 Mark
- (v) the three terms in the sums of squares decomposition

$$SS_T = SS_{\text{Res}} + SS_R :$$

We can compute

$$SS_{\text{Res}} = (n - 2)\hat{\sigma}^2 = 7 \times 1.704^2 = 20.32531.$$

Then,

$$SS_T = \frac{SS_{\text{Res}}}{1 - R^2} = \frac{20.32531}{1 - 0.4948} = 40.23221$$

so therefore

$$SS_R = SS_T - SS_{\text{Res}} = 40.23221 - 20.32531 = 19.9069.$$

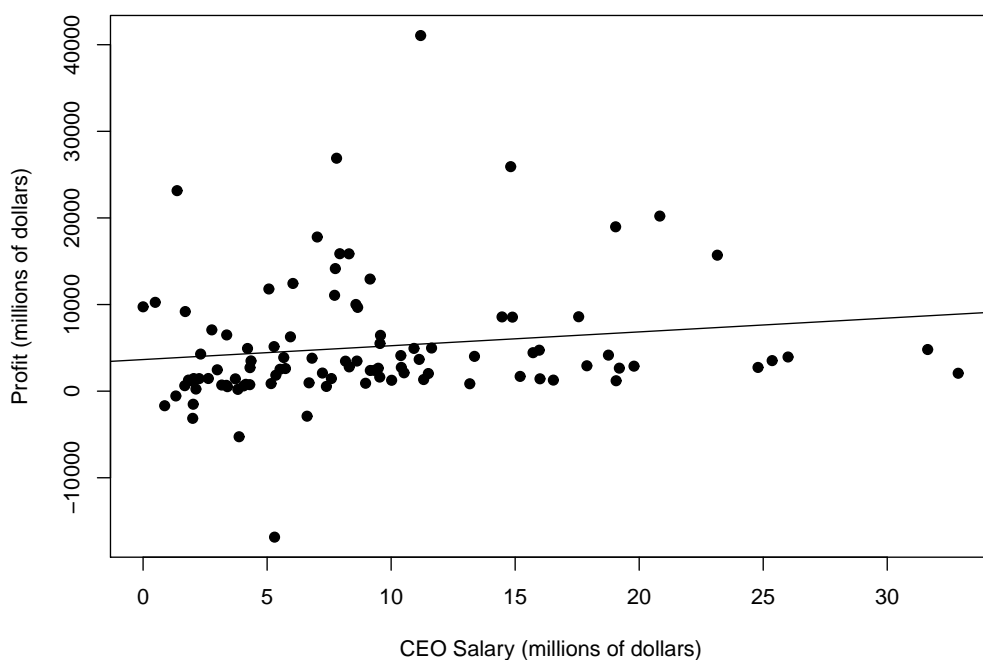
2 Marks

Q2 The following data records Chief Executive Officer (CEO) salary (`CEO.salary`) and annual company profit (`Profit`), each in millions of US dollars, for the most highly paid 100 CEOs in the US for 2012.

The objective of the analysis is to understand whether there is a relationship between annual profit,  $y$ , and CEO salary,  $x$ . An analysis in R is presented below: some values in have been omitted from the output.

```
1 > fit.CEO<-lm(Profit~CEO.salary)
2 > summary(fit.CEO)
3 Coefficients:
4             Estimate Std. Error t value Pr(>|t|)
5 (Intercept)   3648.1     1216.1   3.000  0.00343 **
6 CEO.salary    159.7       105.9   XXXXX  XXXXXXX
7 ---
8 Signif. codes:  0  ***  0.001  **  0.01  *  0.05  .  0.1    1
9
10 Residual standard error: 7301 on 98 degrees of freedom
11 Multiple R-squared:  0.02268,    Adjusted R-squared:  0.01271
12 F-statistic: XXXX on 1 and 98 DF,  p-value: XXXXX
```

The straight line fit is depicted below.



- (a) Is there evidence that the profit of a company increases significantly with CEO salary ? Justify your answer.

Note that the 0.95 and 0.975 quantiles of the Student-t distribution with 98 degrees of freedom are 1.6606 and 1.9845 respectively.

6 Marks

*Question continued on the next page.*