MATH 556: MATHEMATICAL STATISTICS I

SCORE FUNCTION AND FISHER INFORMATION FOR LOCATION-SCALE FAMILIES

The location-scale family for rv X is defined using a linear transformation of a standard variable Z by $X = \mu + \sigma Z$ for $\mu \in \mathbb{R}$ and $\sigma > 0$, and $f_Z(.)$ is a "standard" distribution that may depend on other (fixed) parameters. We restrict attention to the continuous case: we have for the pdf

$$f_X(x; \mu, \sigma) = \frac{1}{\sigma} f_Z\left(\frac{x-\mu}{\sigma}\right).$$

The score function, $S(x; \theta)$, and Fisher information, $I(\theta)$, are defined for $I(m \times 1)$ by

$$\mathbf{S}(x;\theta) = \frac{\partial}{\partial \theta} \left\{ \log f_X(x;\theta) \right\}$$
$$\mathcal{I}(\theta) = \operatorname{Var}_X[\mathbf{S}(X;\theta)] = \mathbb{E}_X \left[\mathbf{S}(X;\theta) \mathbf{S}(X;\theta)^\top \right] - \mathbb{E}_X \left[\mathbf{S}(X;\theta) \right] \mathbb{E}_X \left[\mathbf{S}(X;\theta) \right]^\top.$$

For the location-scale family, we have that m = 2. Under standard regularity conditions we have that

$$\mathbb{E}_{X}[\mathbf{S}(X;\theta)] = \mathbf{0} \qquad \mathcal{I}(\theta) = -\mathbb{E}_{X} \left[\frac{\partial^{2} \log f_{X}(X;\theta)}{\partial \theta \partial \theta^{\top}} \right] = -\mathbb{E}_{X}[\mathbf{\Psi}(X;\theta)].$$

We consider the case where the pdf f_Z has support $\mathbb{Z} = (a,b)$ for values $-\infty \leq a < b \leq \infty$. We have

$$\log f_X(x; \theta) \equiv \log f_X(x; \mu, \sigma) = -\log \sigma + \log f_Z\left(\frac{x - \mu}{\sigma}\right)$$

and so

$$S_{\mu}(x;\mu,\sigma) = \frac{\partial}{\partial \mu} \left\{ \log f_X(x;\mu,\sigma) \right\} = \frac{\partial}{\partial \mu} \left\{ \log f_Z((x-\mu)/\sigma) \right\} = -\frac{1}{\sigma} \frac{\dot{f}_Z((x-\mu)/\sigma)}{f_Z((x-\mu)/\sigma)}$$
(1)

$$S_{\sigma}(x;\mu,\sigma) = \frac{\partial}{\partial \sigma} \left\{ \log f_X(x;\mu,\sigma) \right\} = -\frac{1}{\sigma} + \frac{\partial}{\partial \sigma} \left\{ \log f_Z((x-\mu)/\sigma) \right\} = -\frac{1}{\sigma} - \frac{(x-\mu)}{\sigma^2} \frac{\dot{f}_Z((x-\mu)/\sigma)}{f_Z((x-\mu)/\sigma)}$$
(2)

where $\dot{f}_Z(z) = df_Z(z)/dz$. For the first score function (1):

$$\mathbb{E}_{X} \left[\frac{\dot{f}_{Z}((X-\mu)/\sigma)}{f_{Z}((X-\mu)/\sigma)} \right] = \int_{\mu+\sigma a}^{\mu+\sigma b} \frac{\dot{f}_{Z}((x-\mu)/\sigma)}{f_{Z}((x-\mu)/\sigma)} f_{X}(x;\mu,\sigma) dx = \int_{\mu+\sigma a}^{\mu+\sigma b} \frac{\dot{f}_{Z}((x-\mu)/\sigma)}{f_{Z}((x-\mu)/\sigma)} \frac{1}{\sigma} f_{Z}((x-\mu)/\sigma) dx$$
$$= \int_{a}^{b} \dot{f}_{Z}(z) dz \qquad \text{setting } z = (x-\mu)/\sigma$$
$$= f_{Z}(b) - f_{Z}(a)$$

by standard calculus arguments. Note that this equates to zero if

$$\lim_{z \longrightarrow a} f_Z(z) = \lim_{z \longrightarrow b} f_Z(z) = 0.$$

which certainly holds if the support is the whole of \mathbb{R} .

For the second score function (2): we have that

$$\mathbb{E}_X \left[(X - \mu) \frac{\dot{f}_Z((X - \mu)/\sigma)}{f_Z((X - \mu)/\sigma)} \right] = \int_{\mu + \sigma a}^{\mu + \sigma b} (x - \mu) \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} f_X(x; \mu, \sigma) dx$$

$$= \int_{\mu + \sigma a}^{\mu + \sigma b} (x - \mu) \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} \frac{1}{\sigma} f_Z((x - \mu)/\sigma) dx = \sigma \int_a^b z \dot{f}_Z(z) dz$$

setting $z = (x - \mu)/\sigma$. Using integration by parts

$$\int_{a}^{b} z \dot{f}_{Z}(z) dz = [z f_{Z}(z)]_{a}^{b} - \int_{a}^{b} f_{Z}(z) dz = (b f_{Z}(b) - a f_{Z}(a)) - 1.$$

Note that if $a = -\infty$ and $b = \infty$, this calculation is still valid as

$$\lim_{z \to -\infty} z f_Z(z) = \lim_{z \to \infty} z f_Z(z) = 0$$

because $f_Z(z)$ is integrable, and therefore is $o(|z|^{-(1+\delta)})$ for $\delta > 0$ as $|z| \longrightarrow \infty$. Therefore, from (1) and (2), we have under the usual regularity conditions

$$\mathbb{E}_X\left[\mathbf{S}(X;\mu,\sigma)\right] = -\frac{1}{\sigma} \begin{bmatrix} f_Z(b) - f_Z(a) \\ (bf_Z(b) - af_Z(a)) \end{bmatrix} = \mathbf{0}$$
(3)

whenever $f_Z(a) = f_Z(b) = 0$ which often holds for location-scale models (eg Normal, Cauchy etc).

For the Fisher information, the three distinct elements of $\mathbf{S}(x;\theta)\mathbf{S}(x;\theta)^{\top}$ are

$$\{S_{\mu}(x;\mu,\sigma)\}^{2} = \frac{1}{\sigma^{2}} \left\{ \frac{\dot{f}_{Z}((x-\mu)/\sigma)}{f_{Z}((x-\mu)/\sigma)} \right\}^{2}$$
(4)

$$S_{\mu}(x;\mu,\sigma)S_{\sigma}(x;\mu,\sigma) = \left\{ \frac{1}{\sigma} \frac{\dot{f}_Z((x-\mu)/\sigma)}{f_Z((x-\mu)/\sigma)} \right\} \left\{ \frac{1}{\sigma} + \frac{(x-\mu)}{\sigma^2} \frac{\dot{f}_Z((x-\mu)/\sigma)}{f_Z((x-\mu)/\sigma)} \right\}$$
(5)

$$\{S_{\sigma}(x;\mu,\sigma)\}^{2} = \left\{\frac{1}{\sigma} + \frac{(x-\mu)}{\sigma^{2}} \frac{\dot{f}_{Z}((x-\mu)/\sigma)}{f_{Z}((x-\mu)/\sigma)}\right\}^{2}$$

$$(6)$$

• From (4):

$$\mathbb{E}_{X}\left[\left\{S_{\mu}(X;\mu,\sigma)\right\}^{2}\right] = \frac{1}{\sigma^{2}} \int_{\mu+\sigma a}^{\mu+\sigma b} \left\{\frac{\dot{f}_{Z}((x-\mu)/\sigma)}{f_{Z}((x-\mu)/\sigma)}\right\}^{2} f_{X}(x;\mu,\sigma) dx = \frac{1}{\sigma^{2}} \int_{a}^{b} \frac{\left\{\dot{f}_{Z}(z)\right\}^{2}}{f_{Z}(z)} dz \quad (7)$$

• From (5):

$$\mathbb{E}_X \left[S_{\mu}(x; \mu, \sigma) S_{\sigma}(x; \mu, \sigma) \right]$$

$$= \int_{\mu+\sigma a}^{\mu+\sigma b} \left\{ \frac{1}{\sigma} \frac{\dot{f}_{Z}((x-\mu)/\sigma)}{f_{Z}((x-\mu)/\sigma)} \right\} \left\{ \frac{1}{\sigma} + \frac{(x-\mu)}{\sigma^{2}} \frac{\dot{f}_{Z}((x-\mu)/\sigma)}{f_{Z}((x-\mu)/\sigma)} \right\} f_{X}(x;\mu,\sigma) dx
= \frac{1}{\sigma^{2}} \int_{\mu+\sigma a}^{\mu+\sigma b} \frac{\dot{f}_{Z}((x-\mu)/\sigma)}{f_{Z}((x-\mu)/\sigma)} f_{X}(x;\mu,\sigma) dx + \frac{1}{\sigma^{3}} \int_{\mu+\sigma a}^{\mu+\sigma b} (x-\mu) \left\{ \frac{\dot{f}_{Z}((x-\mu)/\sigma)}{f_{Z}((x-\mu)/\sigma)} \right\}^{2} f_{X}(x;\mu,\sigma) dx
= \frac{1}{\sigma^{2}} \int_{a}^{b} \dot{f}_{Z}(z) dz + \frac{1}{\sigma^{2}} \int_{a}^{b} z \frac{\{\dot{f}_{Z}(z)\}^{2}}{f_{Z}(z)} dz
= \frac{1}{\sigma^{2}} \int_{a}^{b} \frac{\dot{f}_{Z}(z)(z\dot{f}_{Z}(z) + f_{Z}(z))}{f_{Z}(z)} dz$$
(8)

• From (6):

$$\mathbb{E}_{X} \left[\{ S_{\sigma}(X; \mu, \sigma) \}^{2} \right] = \int_{\mu+\sigma a}^{b} \left\{ \frac{1}{\sigma} + \frac{(x-\mu)}{\sigma^{2}} \frac{\dot{f}_{Z}((x-\mu)/\sigma)}{f_{Z}((x-\mu)/\sigma)} \right\}^{2} f_{X}(x; \mu, \sigma) dx$$

$$= \int_{a}^{b} \left\{ \frac{1}{\sigma} + \frac{z}{\sigma} \frac{\dot{f}_{Z}(z)}{f_{Z}(z)} \right\}^{2} f_{Z}(z) dz$$

$$= \frac{1}{\sigma^{2}} \int_{a}^{b} \frac{\{ f_{Z}(z) + z\dot{f}_{Z}(z) \}^{2}}{f_{Z}(z)} dz. \tag{9}$$

Therefore combining (3), (7), (8) and (9) we observe that

$$\mathcal{I}(\theta) \equiv \mathcal{I}(\mu, \sigma) = \frac{1}{\sigma^2} \mathbf{V}_Z$$

where V_Z is a **constant** matrix computed from f_Z , with knowledge of the support (a,b).