1 For t > 0, and constant k > 0

$$P[X \ge t] = P[X + k \ge t + k] \le P[(X + k)^2 \ge (t + k)^2] \le \frac{E_{f_X}[(X + k)^2]}{(t + k)^2}$$

by the Chebychev Lemma. Now if  $k = \sigma^2/t$ , then

$$P[X \ge t] \le \frac{E_{f_X}[(X + \sigma^2/t)^2]}{(t + \sigma^2/t)^2} = \frac{E_{f_X}[(tX + \sigma^2)^2]}{(t^2 + \sigma^2)^2} = \frac{t^2 E_{f_X}[X^2] + 2t E_{f_X}[X] + \sigma^4}{(t^2 + \sigma^2)^2} = \frac{t^2 \sigma^2 + \sigma^4}{(t^2 + \sigma^2)^2}$$

as  $E_{f_X}[X] = 0$  implies  $\sigma^2 = E_{f_X}[X^2]$ . Thus

$$P[X \ge t] \le \frac{\sigma^2}{t^2 + \sigma^2}$$

4	MARKS
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**3** MARKS

2 If  $Y_n = \max\{X_1, \ldots, X_n\}$ , then for  $y \in \mathbb{R}$ ,

(a) For  $Y_n$ 

$$F_{Y_n}(y) = \{F_X(y)\}^n = \left(\frac{1}{2} + \frac{1}{\pi}\arctan(x)\right)^n.$$

But clearly, for any  $y, F_{Y_n}(y) \longrightarrow 0$ , so there is no limiting distribution.

(b) If  $T_n = \pi Y_n/n$ , then

$$F_{T_n}(y) = \{F_X(ny/\pi)\}^n = \left(\frac{1}{2} + \frac{1}{\pi}\arctan(ny/\pi)\right)^n$$

• For *y* < 0,

$$\arctan(ny/\pi) < 0$$
  $\therefore$   $\left(\frac{1}{2} + \frac{1}{\pi}\arctan(ny/\pi)\right) < 1/2$ 

and thus for y < 0,  $F_{T_n}(y) \longrightarrow 0$  as  $n \longrightarrow \infty$ .

• For y = 0,  $\arctan(0) = 0$ , so

$$F_{T_n}(0) = \left(\frac{1}{2}\right)^n \longrightarrow 0$$

and so  $F_{T_n}(y) \longrightarrow 0$  as  $n \longrightarrow \infty$ .

• For *y* > 0,

$$\frac{1}{2} + \frac{1}{\pi}\arctan(ny/\pi) = 1 + \frac{1}{\pi}\arctan(-\pi/(ny)) = 1 + \frac{1}{\pi}\left[-\frac{\pi}{ny} + o(n^{-1})\right]$$

using the approximation given. Thus

Hence, for y > 0, as  $n \longrightarrow \infty$ ,

$$F_{T_n}(y) = \left(1 - \frac{1}{\pi} \left[\frac{\pi}{ny} + o(n^{-1})\right]\right)^n = \left(1 - \frac{1}{ny} + o(n^{-1})\right)^n \longrightarrow \exp\{-1/y\}$$

Note that for arbitrary *A*, *B*,

$$\tan(A+B) = \frac{\sin(A+B)}{\cos(A+B)} = \frac{\sin A \cos B + \cos A \sin B}{\cos A \cos B - \sin A \sin B}$$

so with  $A = \pi/2$ ,  $\cos A = 0$ ,  $\sin A = 1$ , so

$$\tan(\pi/2 + B) = -\frac{\cos B}{\sin B} = -\frac{1}{\tan B}$$

Thus, if  $\tan B = x$ , then

$$\tan(\pi/2 + \arctan(x)) = -\frac{1}{x} \qquad \therefore \qquad \tan(\pi/2 + \arctan(-1/x)) = x$$

and thus

$$\pi/2 + \arctan(-1/x) = \arctan(x)$$

Hence, with  $x = ny/\pi$ ,

$$\frac{1}{2} + \frac{1}{\pi}\arctan(ny/\pi) = 1 + \frac{1}{\pi}\arctan(-\pi/(ny))$$

3 For  $x \in \mathbb{R}$ 

$$f_{X_n}(x) = \frac{1}{\pi} \frac{n}{1+n^2 x^2} \qquad x \in \mathbb{R}.$$

(i) Convergence in *r*th mean to zero;

$$E_{f_{X_n}}[|X_n|^r] = \int_{-\infty}^{\infty} \frac{1}{\pi} \frac{n|x|^r}{1+n^2x^2} = \frac{2}{\pi} \int_0^{\infty} \frac{nx^r}{1+n^2x^2}$$

This integral is **divergent** if  $r \ge 1$ , but **convergent** to zero as  $n \longrightarrow \infty$  if 0 < r < 1. Thus, if we only consider r to take integer values, there is no convergence, but for 0 < r < 1,  $X_n \xrightarrow{r} X$ .

(ii) Convergence in probability; for  $\epsilon > 0$ ,

$$P[|X_n| < \epsilon] = F_{X_n}(\epsilon) - F_{X_n}(-\epsilon) = \frac{1}{\pi} \arctan(n\epsilon) - \frac{1}{\pi} \arctan(-n\epsilon) \longrightarrow 1$$

as  $n \longrightarrow \infty$ . Hence  $X_n \xrightarrow{p} X$ .

Of course

$$X_n \xrightarrow{r} X$$
, some  $r > 0 \implies X_n \xrightarrow{p} X$ 

by general relationships between the modes of convergence.

## MATH 556 ASSIGNMENT 4: SOLUTIONS

**5** Marks

**5** MARKS

4 Suppose  $X_n \xrightarrow{p} 0$ , so that for any  $\epsilon > 0$ ,

$$\lim_{n \longrightarrow \infty} P[|X_n| > \epsilon] = 0$$

Now by iterated expectation, conditioning in turn on the partitioning events

$$(|X_n| \le \epsilon) \qquad (|X_n| > \epsilon)$$

we have

$$\begin{split} E\left[\frac{|X_n|}{1+|X_n|}\right] &= E\left[\frac{|X_n|}{1+|X_n|} \mid |X_n| \le \epsilon\right] P[|X_n| \le \epsilon] + E\left[\frac{|X_n|}{1+|X_n|} \mid |X_n| > \epsilon\right] P[|X_n| > \epsilon] \\ &\leq \frac{\epsilon}{1+\epsilon} \times P[|X_n| \le \epsilon] + 1 \times P[|X_n| > \epsilon] \\ &\longrightarrow \frac{\epsilon}{1+\epsilon} \quad \text{as } n \longrightarrow \infty. \end{split}$$

But this holds for arbitrary  $\epsilon > 0$ , so

$$E\left[\frac{|X_n|}{1+|X_n|}\right] \longrightarrow 0.$$

Conversely, suppose

$$E\left[\frac{|X_n|}{1+|X_n|}\right] \longrightarrow 0.$$

Then, using the Chebychev Lemma and the hint

$$P[|X_n| > \epsilon] = P\left[\frac{|X_n|}{1 + |X_n|} > \frac{\epsilon}{1 + \epsilon}\right] \le \left(\frac{1 + \epsilon}{\epsilon}\right) E\left[\frac{|X_n|}{1 + |X_n|}\right] \longrightarrow 0$$
 as  $n \longrightarrow \infty$ . Thus  
$$P[|X_n| > \epsilon] \longrightarrow 0 \qquad \therefore \qquad X_n \stackrel{p}{\longrightarrow} 0$$

8 MARKS