## MATH 556 - EXERCISES 4

## Not for Assessment.

- 1. Suppose that  $X_1 \sim Geometric(\theta_1)$  and  $X_2 \sim Geometric(\theta_2)$  are independent random variables. Find the pmf of random variable *Y* where  $Y = X_1 + X_2$ .
- 2. Suppose that  $X_1$  and  $X_2$  are random variables with joint pdf given by

$$f_{X_1,X_2}(x_1,x_2) = c|x_1| \exp\left\{-|x_1| - \frac{x_1^2 x_2^2}{2}\right\} \qquad (x_1,x_2) \in \mathbb{R}^2$$

Find the marginal pdf of  $X_1$ , and the conditional pdf of  $X_2$  given  $X_1 = x_1$ , for appropriate values of  $x_1$ . Take care to define the pdfs for all real values of their arguments. Compute the value of constant c

3. The radius of a circle, *R*, is a continuous random variable with density function given by

$$f_R(r) = 6r(1-r) \qquad 0 < r < 1$$

and zero otherwise. Find the joint and marginal pdfs of  $X_1$ , the circumference of the circle, and  $X_2$ , the area of the circle.

4. Suppose that *X* and *Y* are continuous random variables with joint pdf given by

$$f_{X,Y}(x,y) = cx(1-y)$$
  $0 < x < 1, 0 < y < 1$ 

and zero otherwise for some constant *c*. Are *X* and *Y* independent random variables ? Find the value of *c*, and, for the set  $A \equiv \{(x, y) : 0 < x < y < 1\}$ , the probability

$$P_{X,Y}[X < Y] = \iint_A f_{X,Y}(x,y) \, dxdy$$

5. Suppose that *X* and *Y* are continuous random variables with joint pdf given by

$$f_{X,Y}(x,y) = \frac{1}{2x^2y}$$
  $1 \le x < \infty, 1/x \le y \le x$ 

and zero otherwise. Derive

- (i) the marginal pdfs of X and Y
- (ii) the conditional pdf of X given Y = y, and the conditional pdf of Y given X = x.
- (iii) the expectation of Y,  $\mathbb{E}_{Y}[Y]$ .
- 6. Suppose that *X* and *Y* have joint pdf that is constant with support  $\mathbb{X}^{(2)} \equiv (0,1) \times (0,1)$ .
  - (i) Find the marginal pdf of random variables U = X/Y and  $V = -\log(XY)$ , stating clearly the range of the transformed random variable in each case.
  - (ii) Find the pdf and cdf of Z = X Y.

7. (a) Consider random variable *X* with probability function  $P_X$  and cdf  $F_X$ . The indicator random variable for set B,  $\mathbb{1}_B(.)$ , is a transformation of *X*, and is defined by

$$\mathbb{1}_B(X) = \begin{cases} 1 & X \in B \\ 0 & X \notin B \end{cases}$$

Find the pmf/pdf and the expectation of  $\operatorname{rv} \mathbb{1}_B(X)$ .

(b) The expectation of any random variable with pmf/pdf  $f_X$  can be approximated to arbitrary accuracy (under mild conditions) by a *Monte Carlo* simulation procedure: a large sample of simulated values  $x_1, \ldots, x_N$  are generated from  $f_X$ , and then the expectation is approximated by the sample mean to produce the approximation  $\widehat{\mathbb{E}}_X[X]$ .

$$\widehat{\mathbb{E}}_X[X] = \frac{1}{N} \sum_{i=1}^N x_i$$

Using the result in (a), and a Monte Carlo procedure, to approximate the probability

$$P_{\mathbf{X}}[\mathbf{X} \in B]$$

if **X** has a three-dimensional multivariate Normal distribution,  $\mathbf{X} \sim Normal_3(\mathbf{0}, \Sigma)$ , with

$$\Sigma = \begin{bmatrix} 1.0 & 0.2 & -0.5 \\ 0.2 & 2.0 & -0.1 \\ -0.5 & -0.1 & 2.0 \end{bmatrix}$$

and *B* is the set  $\{\mathbf{x} \in \mathbb{R}^3 : x_1^2 + x_2^2 \le x_1 + x_3\}$ . Tabulate the results of five replicate Monte Carlo runs with N = 10000.

R functions for simulating the multivariate Normal distribution include mvrnorm from the MASS library and rmvn from the mvnfast library. Note also that if  $Z_1, \ldots, Z_n$  are independent standard Normal variables, and L is an  $n \times n$  matrix, then

$$\mathbf{Y} = \mathbf{L}\mathbf{Z} \sim Normal(0, \mathbf{L}\mathbf{L}^{\top}).$$

Therefore if **L** is the lower-triangular matrix termed the *Cholesky factor* for  $\Sigma$ , defined by

$$\mathbf{L}\mathbf{L}^{\top} = \Sigma$$

then we can generate X as LZ. The Cholesky factor can be computed in R using the function chol.