

556: MATHEMATICAL STATISTICS I

SCORE FUNCTION AND FISHER INFORMATION FOR LOCATION-SCALE FAMILIES

The location-scale family for rv X is defined using a linear transformation of a standard variable Z by

$$X = \mu + \sigma Z$$

for $\mu \in \mathbb{R}$ and $\sigma > 0$, and $f_Z(\cdot)$ is a “standard” distribution that does not depend on any parameters.

- For a continuous rv, we have for the pdf

$$f_X(x; \mu, \sigma) = \frac{1}{\sigma} f_Z\left(\frac{x - \mu}{\sigma}\right)$$

- For a discrete rv, we have for the pmf

$$f_X(x; \mu, \sigma) = f_Z\left(\frac{x - \mu}{\sigma}\right).$$

In most settings, the discrete location-scale family is not that useful as it effectively amounts merely to a re-labelling of the support of f_Z .

- In both cases, for the cdf, we have

$$F_X(x; \mu, \sigma) = F_Z\left(\frac{x - \mu}{\sigma}\right).$$

The score function, $\mathbf{S}(x; \theta)$, is defined by

$$\mathbf{S}(x; \theta) = \frac{\partial}{\partial \theta} \{\log f_X(x; \theta)\}.$$

If θ is m -dimensional, then $\mathbf{S}(X; \theta)$ is $(m \times 1)$. For the location-scale family, we have that $m = 2$.

For the continuous case, we consider the construction where the pdf f_Z has support $\mathbb{Z} = (a, b)$ for values $-\infty \leq a < b \leq \infty$. We have

$$\log f_X(x; \theta) \equiv \log f_X(x; \mu, \sigma) = -\log \sigma + \log f_Z\left(\frac{x - \mu}{\sigma}\right)$$

and so

$$S_\mu(x; \mu, \sigma) = \frac{\partial}{\partial \mu} \{\log f_X(x; \mu, \sigma)\} = \frac{\partial}{\partial \mu} \{\log f_Z((x - \mu)/\sigma)\} = -\frac{1}{\sigma} \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} \quad (1)$$

$$S_\sigma(x; \mu, \sigma) = \frac{\partial}{\partial \sigma} \{\log f_X(x; \mu, \sigma)\} = -\frac{1}{\sigma} + \frac{\partial}{\partial \sigma} \{\log f_Z((x - \mu)/\sigma)\} = -\frac{1}{\sigma} - \frac{(x - \mu)}{\sigma^2} \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} \quad (2)$$

where

$$\dot{f}_Z(z) = \frac{\partial f_Z(z)}{\partial z}.$$

In the following calculations, integration is over the support (a, b) . For the first score function (1): we have that

$$\begin{aligned}
\mathbb{E}_X \left[\frac{\dot{f}_Z((X - \mu)/\sigma)}{f_Z((X - \mu)/\sigma)} \right] &= \int \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} f_X(x; \mu, \sigma) dx \\
&= \int \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} \frac{1}{\sigma} f_Z((x - \mu)/\sigma) dx \\
&= \int \dot{f}_Z(z) dz && z = (x - \mu)/\sigma \\
&= f_Z(b) - f_Z(a)
\end{aligned}$$

by standard calculus arguments. Note that this equates to zero if

$$\lim_{z \rightarrow a} f_Z(z) = \lim_{z \rightarrow b} f_Z(z) = 0.$$

which certainly holds if the support is the whole of \mathbb{R} .

For the second score function (2): we have that

$$\begin{aligned}
\mathbb{E}_X \left[(X - \mu) \frac{\dot{f}_Z((X - \mu)/\sigma)}{f_Z((X - \mu)/\sigma)} \right] &= \int (x - \mu) \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} f_X(x; \mu, \sigma) dx \\
&= \int (x - \mu) \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} \frac{1}{\sigma} f_Z((x - \mu)/\sigma) dx \\
&= \sigma \int z \dot{f}_Z(z) dz && z = (x - \mu)/\sigma
\end{aligned}$$

and, using integration by parts

$$\int_a^b z \dot{f}_Z(z) dz = [z f_Z(z)]_a^b - \int_a^b f_Z(z) dz = (b f_Z(b) - a f_Z(a)) - 1.$$

Note that if $a = -\infty$ and $b = \infty$, this calculation is still valid as

$$\lim_{z \rightarrow -\infty} z f_Z(z) = \lim_{z \rightarrow \infty} z f_Z(z) = 0$$

because $f_Z(z)$ is integrable, and therefore is $o(|z|^{-(1+\delta)})$ for $\delta > 0$ as $|z| \rightarrow \infty$. Therefore, from (1) and (2), we have under the usual regularity conditions

$$\mathbb{E}_X [\mathbf{S}(X; \mu, \sigma)] = -\frac{1}{\sigma} \begin{bmatrix} f_Z(b) - f_Z(a) \\ (b f_Z(b) - a f_Z(a)) \end{bmatrix}. \quad (3)$$

Notice that this reduces to zero whenever

$$f_Z(a) = f_Z(b) = 0$$

which is a common case when considering location-scale models (eg Normal, Cauchy etc).

The Fisher information, $\mathcal{I}(\theta)$, is then defined as

$$\mathcal{I}(\theta) = \text{Var}_X [\mathbf{S}(X; \theta)] = \mathbb{E}_X \left[\mathbf{S}(X; \theta) \mathbf{S}(X; \theta)^\top \right] - \mathbb{E}_X [\mathbf{S}(X; \theta)] \mathbb{E}_X [\mathbf{S}(X; \theta)]^\top \quad (4)$$

which is an $(m \times m)$ symmetric and non-negative definite matrix. The three distinct elements of $\mathbf{S}(x; \theta)\mathbf{S}(x; \theta)^\top$ are

$$\{S_\mu(x; \mu, \sigma)\}^2 = \frac{1}{\sigma^2} \left\{ \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} \right\}^2 \quad (5)$$

$$S_\mu(x; \mu, \sigma)S_\sigma(x; \mu, \sigma) = \left\{ \frac{1}{\sigma} \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} \right\} \left\{ \frac{1}{\sigma} + \frac{(x - \mu)}{\sigma^2} \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} \right\} \quad (6)$$

$$\{S_\sigma(x; \mu, \sigma)\}^2 = \left\{ \frac{1}{\sigma} + \frac{(x - \mu)}{\sigma^2} \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} \right\}^2 \quad (7)$$

From (5):

$$\begin{aligned} \mathbb{E}_X [\{S_\mu(X; \mu, \sigma)\}^2] &= \frac{1}{\sigma^2} \int \left\{ \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} \right\}^2 f_X(x; \mu, \sigma) dx \\ &= \frac{1}{\sigma^2} \int \frac{\{\dot{f}_Z(z)\}^2}{f_Z(z)} dz \quad z = (x - \mu)/\sigma. \end{aligned} \quad (8)$$

From (6):

$$\begin{aligned} \mathbb{E}_X [S_\mu(x; \mu, \sigma)S_\sigma(x; \mu, \sigma)] &= \int \left\{ \frac{1}{\sigma} \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} \right\} \left\{ \frac{1}{\sigma} + \frac{(x - \mu)}{\sigma^2} \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} \right\} f_X(x; \mu, \sigma) dx \\ &= \frac{1}{\sigma^2} \int \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} f_X(x; \mu, \sigma) dx \\ &\quad + \frac{1}{\sigma^3} \int (x - \mu) \left\{ \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} \right\}^2 f_X(x; \mu, \sigma) dx \\ &= \frac{1}{\sigma^2} \int \dot{f}_Z(z) dz + \frac{1}{\sigma^2} \int z \frac{\{\dot{f}_Z(z)\}^2}{f_Z(z)} dz \quad z = (x - \mu)/\sigma \\ &= \frac{1}{\sigma^2} \int \frac{\dot{f}_Z(z)(z\dot{f}_Z(z) + f_Z(z))}{f_Z(z)} dz \end{aligned} \quad (9)$$

From (7):

$$\begin{aligned} \mathbb{E}_X [\{S_\sigma(X; \mu, \sigma)\}^2] &= \int \left\{ \frac{1}{\sigma} + \frac{(x - \mu)}{\sigma^2} \frac{\dot{f}_Z((x - \mu)/\sigma)}{f_Z((x - \mu)/\sigma)} \right\}^2 f_X(x; \mu, \sigma) dx \\ &= \int \left\{ \frac{1}{\sigma} + \frac{z}{\sigma} \frac{\dot{f}_Z(z)}{f_Z(z)} \right\}^2 f_Z(z) dz \quad z = (x - \mu)/\sigma \\ &= \frac{1}{\sigma^2} \int \frac{\{f_Z(z) + z\dot{f}_Z(z)\}^2}{f_Z(z)} dz. \end{aligned} \quad (10)$$

Therefore combining (3), (8), (9) and (10) we observe from the definition (4) that

$$\mathcal{I}(\theta) \equiv \mathcal{I}(\mu, \sigma) = \frac{1}{\sigma^2} \mathbf{V}_Z$$

where \mathbf{V}_Z is a **constant** matrix computed from f_Z , with knowledge of the support (a, b) .